

Total Points: 60

PSTAT 120B / **FINAL EXAMINATION** / Summer 2024

Instructor: **Ethan P. Marzban**

Name: _____ NetID: _____
(First and Last) (NOT Perm Number)

Your Section: 2pm (Hyuk-Jean) 3pm (Hyuk-Jean) 4pm (Minwoo) 5pm (Minwoo)
(Circle One)

Instructions:

- You will have **2 hours** to complete this exam.
 - Nobody will be permitted to leave the exam room during the last 10 minutes of the exam.
- Please remember to write your name and NetID (not perm number) **at the top of each sheet** of this exam.
- You are allowed the use of **two 8.5 × 11-inch** sheets, front and back, of notes, along with the use of a **calculator**; the use of any and all other resources (including, but not limited to laptops, cell phones, textbooks, etc.) is prohibited.
- Unless otherwise specified, all multiple choice questions have **only one** correct answer.
 - Partial credit is **not** available on multiple choice questions.
- Unless otherwise specified, simplification is not needed; however, all integrals and infinite sums (unless otherwise specified) must be evaluated.
- **Good Luck!!!**

Honor Code: In signing my name below, I certify that all work appearing on this exam is entirely my own and not copied from any external source. I further certify that I have not received any unauthorized aid while taking this exam.

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1 Multiple Choice Questions

Please fill in the bubble(s) on the exam below corresponding to your answer. You do not need to submit any additional work for these questions. No partial credit is available for multiple choice questions.

- (1 point) **True or False:** Given a sample $\vec{Y} = \{Y_i\}_{i=1}^n$ from a distribution with unknown parameter θ , the entire sample \vec{Y} is a sufficient statistic for θ .
 True False
- (1 point) Consider $Y_1, \dots, Y_n \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mu, 1)$, where $\mu \in \mathbb{R}$ is unknown. Is it possible to find an estimator $\hat{\mu}_n$ for μ whose variance is smaller than the Cramér-Rao Lower Bound? (Read the problem statement carefully!)
 Yes, it is possible No, it is not possible
- (1 point) Consider an i.i.d. sample $\vec{Y} := \{Y_i\}_{i=1}^n$ from a distribution with unknown parameter θ . Suppose that $U := \sum_{i=1}^n \sqrt{Y_i}$ is a sufficient statistic for θ . Which of the following must also be a sufficient statistic for θ ? (Remember, there is only one correct answer.)
 $\sum_{i=1}^n Y_i$
 $2 \sum_{i=1}^n \sqrt{Y_i}$
 $\frac{1}{n} \sum_{i=1}^n Y_i$
 Y_1
 None of the above
- (1 point) Which of the following properties does the maximum likelihood estimator **not** possess in general?
 Asymptotic normality
 Asymptotic efficiency
 Consistency
 Unbiasedness (for any sample size)
 None of the above
- (1 point) If $(Y \mid X = x) \sim \text{Exp}(x)$ and $X \sim \text{Gamma}(3, 2)$, what is the correct value of $\mathbb{E}[Y]$?
 1.500
 6.000
 x [this is a lowercase x]
 X [this is an uppercase X]
 None of the above

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2 Free-Response Questions

1. (6 points) Let $Y_1, Y_2 \stackrel{\text{i.i.d.}}{\sim} \text{Exp}(\theta)$, and define $U := Y_2/Y_1$. Show that $F_U(u)$, the CDF (cumulative distribution function) of U , is given by

$$F_U(u) = \begin{cases} 0 & \text{if } u < 0 \\ \frac{u}{u+1} & \text{if } u \geq 0 \end{cases}$$

For full credit, you must set up and evaluate a double integral corresponding to $F_U(u)$, and you **must sketch the region of integration.**

2. Let $Y_1, \dots, Y_n \stackrel{\text{i.i.d.}}{\sim} f(y; \theta)$ where

$$f(y; \theta) = \theta y^{\theta-1} \cdot \mathbb{1}_{\{0 < y < 1\}}$$

and $\theta > 0$ is an unknown parameter.

(a) (3 points) Show that $U := \left(\prod_{i=1}^n Y_i \right)$ is a sufficient statistic for θ .

(b) (3 points) Let $W_i := -\ln(Y_i)$. Use any of the methods discussed in this course to show that $W_i \sim \text{Exp}(1/\theta)$.

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(c) (3 points) Show that

$$V := \left(2\theta \sum_{i=1}^n [-\ln(Y_i)] \right) \sim \chi_{2n}^2$$

You may use any result from class, however you must clearly state which result/s you are using. Don't skip any steps!

3. Let $Y_1, \dots, Y_n \stackrel{\text{i.i.d.}}{\sim} f(y; \beta)$ where

$$f(y; \beta) = \frac{3\beta^3}{y^4} \cdot \mathbf{1}_{\{y \geq \beta\}}$$

and $\beta > 0$ is an unknown constant. A fact you may use without proof is that $\mathbb{E}[Y_1] = 3\beta/2$.

(a) (2 points) Derive an expression for $\hat{\beta}_{\text{MM}}$, the method of moments estimator for β .

(b) (2 points) Show that $\hat{\beta}_{MM}$ is a consistent estimator for β - don't just cite the property that method of moments estimators are typically consistent!

(c) (3 points) Derive an expression for $\mathcal{L}_{\vec{Y}}(\beta)$, the likelihood of the sample $\vec{Y} = \{Y_i\}_{i=1}^n$.

(d) (2 points) Use your answer to part (c) to find $\hat{\beta}_{MLE}$, the method of moments estimator for β .

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4. Let $Y_1, \dots, Y_n \stackrel{\text{i.i.d.}}{\sim} f(y; \theta)$ where

$$f(y; \theta) = \frac{3y^2}{\theta^3} \cdot \mathbb{1}_{\{0 \leq y \leq \theta\}}$$

and $\theta > 0$ is an unknown parameter. A fact you may use, without proof, is that the CDF (cumulative distribution function) of Y_1 is given by

$$F_{Y_1}(y) = \begin{cases} 0 & \text{if } y < 0 \\ (y/\theta)^3 & \text{if } 0 \leq y < \theta \\ 1 & \text{if } y \geq \theta \end{cases}$$

Additionally, define $U := Y_{(n)}/\theta$, where $Y_{(n)} := \max_{1 \leq i \leq n} \{Y_i\}$ denotes the sample maximum (i.e. the n^{th} order statistic).

(a) (3 points) Show that the density of $Y_{(n)}$, the n^{th} order statistic, is given by

$$f_{Y_{(n)}}(y) = \frac{3ny^{3n-1}}{\theta^{3n}} \cdot \mathbb{1}_{\{0 \leq y \leq \theta\}}$$

(b) (3 points) Show that U has density

$$f_U(u) = 3nu^{3n-1} \cdot \mathbb{1}_{\{0 \leq u \leq 1\}}$$

and use the definition of a pivotal quantity to argue that U is a pivotal quantity for θ .

(c) (4 points) Find constants a and b such that $\mathbb{P}(U < a) = \alpha/2$ and $\mathbb{P}(U > b) = \alpha/2$.

(d) (4 points) Combine your answers to parts (a) and (b) to construct a $(1 - \alpha) \times 100\%$ confidence interval for θ .

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5. As she is watching the 2024 Paris Olympics, Bijonka becomes interested in performing inference on the true average length (in minutes) of a tennis match. Among an i.i.d. sample of 25 tennis matches, she observes a sample average length of 140 minutes and a sample standard deviation of 10 minutes. Suppose she wishes to test the claim that the true average length of a tennis match is 150 minutes against a two-sided alternative, using a 5% level of significance. Assume the lengths of tennis matches are normally distributed.

(a) (1 point) Let μ denote the true average length (in minutes) of a tennis match. State the null and alternative hypotheses.

(b) (1 point) Should Bijonka perform a Z -test or a T -test? Justify your answer.

(c) (2 points) Compute the observed value of the test statistic.

(d) (2 points) Derive an expression for the p -value of the observed value of the test statistic. You may leave your answer in terms of the CDF of a distribution.

(e) (2 points) Suppose that the p -value associated with Biyonka's data is 0.01 (which is **not** to say this is the right answer to part (c) above!) State the conclusions of the test, and phrase your findings in the context of the problem.

6. Let $(Y_1, Y_2) \sim f_{Y_1, Y_2}(y_1, y_2) = 2 \cdot \mathbb{1}_{\{0 \leq y_1 < y_2 \leq 1\}}$.

(a) (3 points) Derive an expression for $f_{Y_1}(y_1)$, the marginal density of Y_1 . Be sure to include the support of Y_1 as well!

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- (b) (3 points) Derive an expression for $f_{Y_1|Y_2}(y_1 | y_2)$, the conditional density of $(Y_1 | Y_2 = y_2)$. Be sure to not only include the support, but also the values of y_2 for which the density is defined! A fact you may use without proof is that the marginal density of Y_2 is given by

$$f_{Y_2}(y_2) = 2y_2 \cdot \mathbf{1}_{\{0 \leq y_2 \leq 1\}}$$

- (c) (3 points) Compute $\mathbb{E}[Y_1 | Y_2 = y_2]$.

You may use the remainder of this page for scratch work; please note that nothing written on this page will be graded.

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