DISCUSSION WORKSHEET 05

PSTAT 120B: Mathematical Statistics, I **Summer Session A, 2024** with Instructor: Ethan P. Marzban

Conceptual Review

- (a) What is a **population moment**? What is a **sample moment**?
- (b) What is the **Method of Moments**? Does the Method of Moments produce consistent estimators? What about unbiased estimators?
- (c) What is a **likelihood**?

Problem 1

Let $Y_1,\cdots,Y_n \stackrel{\mathsf{i.i.d.}}{\sim} f(y;\alpha)$ where

$$
f(y;\alpha)=\frac{1+\alpha y}{2}\cdot 1\!\!1_{\{-1\leq y\leq 1\}}
$$

where $\alpha \in [-1, 1]$ is an unknown parameter.

(a) Find $\widehat{\alpha}_{MM}$, the method of moments estimator for α .

Solution: We first compute

$$
\mathbb{E}[Y_1] = \int_{-1}^1 y \cdot \frac{1 + \alpha y}{2} \, \mathrm{d}y = \frac{\alpha}{3}
$$

Therefore, $\widehat{\alpha}_{\mathsf{MM}}$ must satisfy

$$
\frac{\widehat{\alpha}_{\text{MM}}}{3} = \overline{Y}_n \iff \frac{\widehat{\alpha}_{\text{MM}}}{3} = 3\overline{Y}_n
$$

(b) Show that $\widehat{\alpha}_{MM}$ is a consistent estimator for α .

Solution:

- By the WLLN, $\overline{Y}_n \overset{\mathbb{P}}{\longrightarrow} \mathbb{E}[Y_1] = \frac{\alpha}{3}$
- By the CMT with $g(x)=3x$, we have $3\overline{Y}_n\stackrel{\mathbb{P}}{\longrightarrow}3\mathbb{E}[Y_1]=e\cdot\frac{\alpha}{3}=\alpha$ \checkmark
- (c) Suppose we obtain the following observed instance of a sample from the above-stated distribution:

 $\vec{y} = (-0.38, 0.61, -0.13, 0.79, 0.89, -0.90, 0.11, 0.80)$

What is an appropriate estimate for α based on this sample, using the method of moments?

Solution: We simply need to plug into our estimator to obtain an estimate:

$$
\overline{y} = \frac{1}{8}(-0.38 + 0.61 - 0.13 + \dots + 0.80) = 0.22375
$$

$$
3\overline{y} = 3 \cdot 0.22375 \approx 0.67125
$$

(d) What is the approximate sampling distribution of $\widehat{\alpha}_{\mathsf{MM}}$, assuming a large sample size?

Solution: Assuming a large sample size, the CLT tells us that

$$
\frac{\sqrt{n}(\overline{Y}_n-\mu)}{\sigma}\stackrel{d}{\approx}\mathcal{N}(0,1)
$$

Another way to say this is, for large n ,

$$
\overline{Y}_n \stackrel{d}{\approx} \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)
$$

meaning

$$
3\overline{Y}_n \stackrel{d}{\approx} \mathcal{N}\left(3\mu\,,\,\frac{9\sigma^2}{n}\right)
$$

We already computed $\mu=\alpha/3$ in part (a); now all we need is to compute σ^2 :

$$
\mathbb{E}[Y_i^2] = \int_{-1}^1 y^2 \cdot \frac{1 + \alpha y}{2} dy dy = \frac{1}{3}
$$

Var(Y_i) = $\mathbb{E}[Y_i^2] - (\mathbb{E}[Y_i])^2 = \frac{1}{3} - (\frac{\alpha}{3})^2 = \frac{3 - \alpha^2}{9}$

Hence,

$$
\widehat{\alpha}_{\text{MM}} = 3\overline{Y}_n \stackrel{d}{\approx} \mathcal{N}\left(\alpha, \frac{3-\alpha^2}{n}\right)
$$

(e) Use your answer to part (d) to approximate the probability that $\widehat{\alpha}_{\sf MM}$ exceeds 1. Why is this a problem (i.e. why would we want this probability to be as small as possible)?

Solution:

$$
\mathbb{P}(\widehat{\alpha}_{\text{MM}} > 1) = \mathbb{P}\left(\frac{\widehat{\alpha}_{\text{MM}} - \alpha}{\sqrt{\frac{3 - \alpha^2}{n}}} > \frac{1 - \alpha}{\sqrt{\frac{3 - \alpha^2}{n}}}\right)
$$

$$
\approx 1 - \Phi\left(\frac{1 - \alpha}{\sqrt{\frac{3 - \alpha^2}{n}}}\right) = \frac{1 - \Phi\left(\sqrt{n} \cdot \frac{1 - \alpha}{\sqrt{3 - \alpha^2}}\right)}{\sqrt{\frac{3 - \alpha^2}{n}}}
$$

The problem explicitly tells us that α is constrained to be between -1 and 1 (indeed, otherwise

the density isn't even a valid probability density!). Hence, obtaining an estimate above 1 is problematic. However, note that as $n \to \infty$ this probability goes to 0, so in the long run we should be okay! (Also, keep in mind that the CLT provides an *approximation* to the sampling distribution of $\widehat{\alpha}_{\sf MM}$.)